## Change-point detection for Piecewise Deterministic Markov Processes

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Motivation: Stochastic control

## Outline

#### Motivation: Stochastic control Dynamic optimization Examples Piecewise deterministic Markov Processes Impulse control

Change-point detection problem

Numerical approximation

Simulation study

Conclusion and perspectives

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# Stochastic control problems Definition



#### Dynamic decision making problems

act on a time-dependent process to change its dynamics



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  - continuously: use the accelerator pedal in a car
  - punctually: change gear

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  - other cars
  - unknown route

#### Stochastic control problems Questions of interest

Dynamic decision making problems under uncertainty

value function: best mean performance

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[Pasin 18]

Population: HIV patients



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  - dose
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- Possible actions: cycles of injections of IL
  - number of injections
  - dose
  - dates of injection
- Objective: minimize the time spent with low CD4<sup>+</sup> T lymphocytes count
- Sources of randomness
  - random response to injections
  - individual variability between patients

#### [Pasin 18]

Examples of optimally controlled CD4<sup>+</sup> T trajectories



[Geeraert 17]

Object of interrest: multi-component optronic camera



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- Possible actions: maintenance
  - repair or replace
  - which components
  - dates of intervention
- Objective: minimize the unavailability + maintenance cost
- Sources of randomness
  - random degradation or failure times for each component

### [Geeraert 17]

#### Reference policy

- send camera to the workshop one day after failure or deterioration
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#### Minimal cost (value function)

- maintenance authorized only after failure or deterioration: 20% lower
- maintenance authorized at all times: 38% lower

### Common points

- family of underlying stochastic models PDMPs
- type of optimization problem: impulse control

#### Davis (80's)

General class of non-diffusion dynamic stochastic hybrid models: deterministic motion punctuated by random jumps.

Starting point

$$X_0 = (m, x)$$



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General class of non-diffusion dynamic stochastic hybrid models: deterministic motion punctuated by random jumps.

 $X_t$  follows the deterministic flow until the first jump time  $T_1 = S_1$ 

$$X_t = ig(m, \phi_m(x, t)ig), \quad \mathbb{P}_{(m, x)}(S_1 > t) = \mathrm{e}^{-\int_0^t \lambda_m ig(\phi_m(x, s)ig) ds}$$



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Post-jump location  $(m_1, x_{T_1})$  selected by the Markov kernel

 $Q_m(\phi_m(x, T_1), \cdot)$ 



#### Davis (80's)

General class of non-diffusion dynamic stochastic hybrid models: deterministic motion punctuated by random jumps.

 $X_t$  follows the flow until the next jump time  $T_2 = T_1 + S_2$ 

$$X_{T_1+t} = (m_1, \phi_{m_1}(x_{T_1}, t)), \quad t < S_2$$



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Post-jump location  $(m_2, x_{T_2})$  selected by Markov kernel

 $Q_{m_1}(\phi_{m_1}(x_{T_1},S_2),\cdot)\ldots$ 



## Applications

#### Applications of PDMPs

Engineering systems, operations research, management science, economics, internet traffic, dependability and safety, neurosciences, biology, ...

- mode: nominal, failures, breakdown, environment, number of individuals, response to a treatment, ...
- Euclidean variable: pressure, temperature, time, size, potential, protein level, ...

## Impulse control problem

#### Impulse control

Select

- intervention dates
- new starting point for the process at interventions

to minimize a cost function

- repair a component before breakdown
- change treatment before relapse

...

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  - heuristics, no mathematical proof

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#### In all cases, the process is perfectly observed at all times CIMOM Mayotte November 2018

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### If the jump times are not observed?

Jump times can be

- date when CD4<sup>+</sup> T count reach 500 threshold
- random failure/deterioration dates

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- ▶ [BdSD 12] Optimal stopping
  - jump times observed
  - post-jump locations observed through noise

Numerical approximation of the value function and  $\epsilon\text{-optimal}$  stopping time

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Optimality equation, existence of optimal policies

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Optimality equation, existence of optimal policies

### No information on the jump times $\Rightarrow$ very difficult problem

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## Change-point detection

#### Simplest special case

- only one jump of the mode variable
- discrete noisy observations of the continuous variable on a regular time grid

#### Optimal stopping = Change-point detection

Aim: numerical approximation to

- detect the change-point at best (not too early/late)
- estimate the new mode after the jump

Population: cancer patients



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- Possible actions: change treatment
  - treatment 1
  - treatment 2
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- Possible actions: change treatment
  - treatment 1
  - treatment 2
  - dates of change
- Objective: maximize life time of the patient with minimal secondary effect
- Sources of randomness
  - relapse date
  - relapse type
- Observations: cancer cell loads (or proxy) at some regularly spaced measurement times, e.g. every 3 month

## Outline

Motivation: Stochastic control

#### Change-point detection problem PDMP model State of the art on change-point detection MDP model

Numerical approximation

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## Simple PDMP model

- State space *E* × ℝ = {0, 1, ..., *d*} × ℝ × ℝ: mode, position, time
- Starting point  $X_0 = (0, x, 0)$ , flow  $\Phi_0$
- time-dependent Jump intensity  $\lambda_0(x, u) = \lambda(u)$
- Jump kernel: position and time continuous, switch to mode i with probability p<sub>i</sub>



### Observations

- Observation times  $t_n = \delta n$
- Noisy observations of the positions  $Y_n = F(x_{t_n}) + \epsilon_n$



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## Example: flat/exponential model

► d = 3 possible post-jump modes, same probability p<sub>i</sub> = 1/3, starting from x<sub>0</sub> = 1

• 
$$\Phi_0(x,t) = x, \ \Phi_1(x,t) = xe^{0.1t}, \ \Phi_2(x,t) = xe^{0.5t}, \ \Phi_3(x,t) = xe^{1t}$$



## Segmentation

- data collected until the time horizon
- a posteriori reconstruction of the change point

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Irrelevant in our medical context: change must be detected as soon as possible

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### Moving average

- compute the average of past data over a moving window
- detect rupture when the average exceeds some threshold

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- compute the average of past data over a moving window
- detect rupture when the average exceeds some threshold

Works well if

- data are centered before the rupture
- date have a positive trend after the rupture
- data have low variance
- small time interval between data



## Kalman Filter

- discrete-time linear gaussian model observed throuh gaussian additive noise
- best mean squares approximation of the hidden variable given the observations
- small time interval between data



## State of the art on change point detection

No generic method available if

- long interval between 2 observations
- non gaussian-linear model
- non additive noise
- aim is to detect rupture and new mode after rupture

Finite horizon 
δN

- Finite horizon  $\delta N$
- Admissible stopping times  $\tau$ :  $\mathcal{F}^{Y}$ -measurable
- Admissible decisions A:  $\{0, 1, \dots, d\}$  valued,  $\mathcal{F}_{\tau}^{Y}$ -measurable

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- Admissible decisions A:  $\{0, 1, \dots, d\}$  valued,  $\mathcal{F}_{\tau}^{Y}$ -measurable
- Cost per stage before stopping
  - c(0, x, y) = 0 rightfully not stopped
  - $c(m \neq 0, x, y) = \beta_i \delta$  lateness penalty
- Terminal cost at stopping
  - C(m, x, y, 0) = c(m, x, y) no stopping before the horizon
  - $C(0, x, y, a \neq 0) = \alpha$  early stopping penalty
  - $C(m \neq 0, x, y, a = m) = 0$  good mode selection
  - $C(m \neq 0, x, y, a \neq 0, m) = \gamma$  wrong mode penalty

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Cost of admissible strategy  $(\tau, A)$ 

$$J(\tau, A, (m, x, y)) = \mathbb{E}_{(m, x, y)} \left[ \sum_{n=0}^{(\tau-1) \wedge N} c(X_n, Y_n) + C(X_{\tau \wedge N}, Y_{\tau \wedge N}, A) \right]$$

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- ► Filter process  $\Theta_n(A \times B) = \mathbb{P}_{(0,x,y)}(X_{\delta n} \in A \times B | \mathcal{F}_n^Y)$
- ► (Θ<sub>n</sub>, Y<sub>n</sub>) time inhomogeneous Markov chain with explicit transition kernels R'<sub>n</sub> on P(E) × ℝ
- ► cost functions  $c'(\theta, y) = \int c(m, x, y) d\theta(m, x)$ ,  $C'(\theta, y, a) = \int C(m, x, y, a) d\theta(m, x)$

#### Fully observed optimal stopping problem

Minimize over all admissible strategies  $(\tau, a)$ 

$$J'(\tau, A, (\theta, y)) = \mathbb{E}_{(\theta, y)} \left[ \sum_{n=0}^{(\tau-1) \wedge N} c'(\Theta_n, Y_n) + C'(\Theta_{\tau \wedge N}, Y_{\tau \wedge N}, A) \right]$$

### Aim

- numerical approximation of the value function
- computable (optimal ?) strategy

#### Difficulties

- measure-valued filter process: recursive equations but not simulatable
- curse of dimensionality

## Outline

Motivation: Stochastic control

Change-point detection problem

Numerical approximation Approach Optimal quantization Convergence results Computable strategy

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Approach

# Dynamic programming

#### Value function

$$\mathcal{L}'(\theta, y) = \inf_{(\tau, A)} J'(\tau, A, (\theta, y))$$
  
=  $\inf_{(\tau, A)} \mathbb{E}_{(\theta, y)} \left[ \sum_{n=0}^{(\tau-1) \wedge N} c'(\Theta_n, Y_n) + C'(\Theta_{\tau \wedge N}, Y_{\tau \wedge N}, A) \right]$ 

#### Dynamic programming

$$\begin{aligned} v'_{N}(\theta, y) &= \min_{0 \le a \le d} C'(\theta, y, a) \\ v'_{k}(\theta, y) &= \min \left\{ \min_{1 \le a \le d} C'(\theta, y, a); c'(\theta, y) + R'_{k} v'_{k+1}(\theta, y) \right\} \\ v'_{0} &= V' \end{aligned}$$

## Approach

 Discretize the kernels R'<sub>k</sub> to discretize the Dynamic programming operators

based on simulation-based discretization grids of the chain  $(\Theta_n, Y_n)$ .



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based on simulation-based discretization grids of the chain  $(\Theta_n, Y_n)$ .

#### Problems

•  $\Theta_n$  is not simulatable

$$\Theta_{n+1}(A) = \frac{\int_{\mathbb{X}} P_n(H_{Y_{n+1}} \mathbb{1}_A)(m, x) d\Theta_n(m, x)}{\int_{\mathbb{X}} P_n(H_{Y_{n+1}})(m, x) d\Theta_n(m, x)}$$

approximation in 2 steps: approximate simulation of Θ<sub>n</sub> + discretization of the approximation

#### Discretization

$$\begin{aligned} X_t &= (m_y, x_t, t) \\ E \times \mathbb{R}, \ P \end{aligned}$$



#### Discretization

$$X_{t} = (m_{y}, x_{t}, t)$$

$$E \times \mathbb{R}, P$$

$$\downarrow$$

$$X_{n} = (m_{t_{n}}, x_{t_{n}})$$

$$E, P_{n}$$
observations
$$\downarrow Y_{n} = F(X_{n}) + \varepsilon$$

$$(X_{n}, Y_{n})$$

$$E \times \mathbb{R}, R_{n}$$

#### Discretization

$$\begin{aligned} X_t &= (m_y, x_t, t) \\ E \times \mathbb{R}, P \\ & \downarrow \\ X_n &= (m_{t_n}, x_{t_n}) \\ E, P_n \\ observations & \bigvee Y_n &= F(X_n) + a \\ (X_n, Y_n) \\ E \times \mathbb{R}, R_n \\ filtering & \downarrow \Psi \\ (\Theta_n, Y_n) \\ \mathcal{P}(E) \times \mathbb{R}, R'_n \\ programming \\ & \downarrow \\ v'_n(\Theta_n, Y_n) \end{aligned}$$
$$\begin{array}{c} X_{t} = (m_{y}, x_{t}, t) \\ E \times \mathbb{R}, P \\ \downarrow \\ X_{n} = (m_{t_{n}}, x_{t_{n}}) \quad (\bar{m}_{t_{n}}, \bar{x}_{t_{n}}) = \bar{X}_{n} \\ E, P_{n} \xrightarrow{quantization} \Omega_{n}, \bar{P}_{n} \\ observations \\ Y_{n} = F(X_{n}) + \varepsilon_{n} \\ (X_{n}, Y_{n}) \\ E \times \mathbb{R}, R_{n} \\ filtering \\ \psi \\ (\Theta_{n}, Y_{n}) \\ \mathcal{P}(E) \times \mathbb{R}, R'_{n} \\ programming \\ v'_{n}(\Theta_{n}, Y_{n}) \end{array}$$

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# Quantization

[P 98], [PPP 04], [PRS05], ...

Quantization of a random variable  $X \in L^2(\mathbb{R}^q)$ 

Approximate X by  $\hat{X}$  taking finitely many values such that  $||X - \hat{X}||_2$  is minimum

- Find a finite weighted grid  $\Gamma$  with  $|\Gamma| = K$
- Set  $\widehat{X} = p_{\Gamma}(X)$  closest neighbor projection

#### Asymptotic properties

If  $E[|X|^{2+\eta}] < +\infty$  for some  $\eta > 0$  then

$$\lim_{K \to \infty} K^{1/q} \min_{|\Gamma| \le K} \|X - \widehat{X}^{\Gamma}\|_2 = C$$

# Algorithms

There exist algorithms providing

- ► F
- ▶ law of  $\widehat{X}$

► transition probabilities for quantization of Markov chains Example:  $\mathcal{N}(0, l_2)$ :



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- ► law of  $\widehat{X}$

• transition probabilities for quantization of Markov chains Example:  $\mathcal{N}(0, I_2)$ :

















# Assets and drawbacks of quantization

#### Assets

- a simulator of the target law is enough to build the grids
- automatic construction of grids
- convergence rate for  $\mathbb{E}[|f(X) f(\widehat{X})|]$  if f lipschitz
- empirical error measure by Monte Carlo

#### Drawbacks

- computation time
- curse of dimension
- open questions of convergence of the algorithms

## Convergence

#### Technical assumptions

$$\begin{aligned} |v_0'(\delta_{(0,x_0)},y_0)-\bar{v}_0'(\delta_{(0,x_0)},y_0)| &\leq \sum_{n=0}^{N-1} a_n \mathbb{E}[|\bar{X}_n-X_n|] \\ &= O(N_{\Omega}^{-1}) \end{aligned}$$

$$\begin{aligned} |\hat{v}_0'(\delta_{(0,x_0)}, y_0) - \bar{v}_0'(\delta_{(0,x_0)}, y_0)| \\ &\leq \sum_{n=0}^N c_n \left( \mathbb{E}\left[ \left| \hat{Y}_n - \bar{Y}_n \right| \right] + \mathbb{E}\left[ \|\hat{\Theta}_n - \bar{\Theta}_n\|_{n,1} \right] \right) \\ &= O(N_{\Gamma}^{-1/N_{\Omega}}) \end{aligned}$$

# Candidate computable strategy

#### Dynamic programming

$$\hat{v}'_{N}(\hat{\theta}, \hat{y}) = \min_{0 \le a \le d} C'(\hat{\theta}, \hat{y}, a)$$

$$\hat{v}'_{k}(\hat{\theta}, \hat{y}) = \min\left\{\min_{1 \le a \le d} C'(\hat{\theta}, \hat{y}, a); c'(\hat{\theta}, \hat{y}) + \hat{R}'_{k} \hat{v}'_{k+1}(\hat{\theta}, \hat{y})\right\}$$

#### Set

► 
$$r_N(\cdot) = 0$$
,  $a_N(\cdot) = 0$  if  $\hat{v}'_N(proj_{\Gamma_N}(\cdot)) = C'(proj_{\Gamma_N}(\cdot), 0)$   
►  $r_N(\cdot) = 1$ ,  $a_N(\cdot) = i$  if  $\hat{v}'_N(proj_{\Gamma_N}(\cdot)) = C'(proj_{\Gamma_N}(\cdot), i)$ 

# Candidate computable strategy

#### Dynamic programming

$$\hat{v}_{N}'(\hat{\theta}, \hat{y}) = \min_{0 \le a \le d} C'(\hat{\theta}, \hat{y}, a)$$

$$\hat{v}_{k}'(\hat{\theta}, \hat{y}) = \min\left\{\min_{1 \le a \le d} C'(\hat{\theta}, \hat{y}, a); c'(\hat{\theta}, \hat{y}) + \hat{R}_{k}' \hat{v}_{k+1}'(\hat{\theta}, \hat{y})\right\}$$

#### Set

CIMOM





















Simulation study

## Outline

Motivation: Stochastic control

Change-point detection problem

Numerical approximation

Simulation study Linear model Non linear model

Conclusion and perspectives



► 
$$d = 3$$
,  $p_i = 1/3$ ,  $x_0 = 1$   
►  $\Phi_0(x, t) = x$ ,  $\Phi_1(x, t) = xe^{0.1t}$ ,  $\Phi_2(x, t) = xe^{0.5t}$ ,  
 $\Phi_3(x, t) = xe^{1t}$ 

 $\blacktriangleright~\beta=1$  (late detection),  $\gamma=1.5$  (wrong mode),  $\delta=1/6$ 



CIMOM

Mayotte

November 2018

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#### Linear model

# Flat/exponential model

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$$d = 3$$
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 $\Phi_3(x, t) = xe^{1t}$ 

• 
$$eta=1$$
 (late detection),  $\gamma=1.5$  (wrong mode),  $\delta=1/6$ 

	MA	KF	PDMP
linear link function $F(x) = x$	1.42	1.60	1.00
inverse link function $F(x) = 1/x$	2.17	1.81	1.17

# Non-linear model

$$b d = 1, x_0 = (0,0) b \Phi_0((x,u),t) = (sin(3\pi(u+t)), u+t), \Phi_1((x,u),t) = (sin(5\pi(u+t)), u+t)$$

• 
$$\delta = 1/6$$
, noise variance 1



#### Non linear model

# Non-linear model

► 
$$d = 1, x_0 = (0, 0)$$
  
►  $\Phi_0((x, u), t) = (sin(3\pi(u + t)), u + t),$   
 $\Phi_1((x, u), t) = (sin(5\pi(u + t)), u + t)$ 

•  $\delta = 1/6$ , noise variance 1



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- Change-point detection method for continuous-time jump dynamics, able to detect a jump and select the post-jump mode
- For general flows but dimension 1

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#### To be done

- Real data applications
- Theoretical validity of the stopping rule
- Allow to stop between observations

- Change-point detection method for continuous-time jump dynamics, able to detect a jump and select the post-jump mode
- For general flows but dimension 1

#### To be done

- Real data applications
- Theoretical validity of the stopping rule
- Allow to stop between observations
- Several jumps and detections
- Impulse control: select an action that changes the dynamics
- Optimally decide the next observation date

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